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The Performance Characteristics of Some Reliability  
Growth Models

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## SUMMARY

A reliability growth model is an analytical model that accounts for changes in reliability due to design changes and other corrective actions taken during the development and testing phases of a reliability program. This paper describes the results of a Monte Carlo study comparing the performance characteristics of four reliability growth models that have been proposed in the reliability literature.

## 1. INTRODUCTION

A reliability growth model is an analytical model that accounts for changes in reliability (usually an improvement) due to design changes and other corrective actions taken during the development and testing phases of a reliability program. A good reliability growth model would be useful in determining a test plan that will lead to the development of an acceptable product within a reasonable time period. Once a test plan is developed the model may be used to monitor the progress of the reliability program and if warranted, initiate changes such as accelerated testing. At any point in time during the test phase an idea as to how much additional testing would be required to achieve a reliability goal may be obtained using the model.

There are several reliability growth models available in the literature. Some of these models are probabilistic in nature i.e., the models do not allow for incorporation of information obtained during the test phase. The second type of models, generally known as statistical models, involve unknown parameters that can be estimated using available test data. This paper describes the results of a study comparing the performance for four statistical models. Two of these models are intended for use with time to failure data. The other two models use bernoulli type data; a "success" is taken to mean "the satisfactory operation for a specified period of time" and the specified time period is often referred to as "mission time".

A description of the reliability growth models selected for this study is given in Section 2. The Monte Carlo procedure used is described in Section 3 and Section 4 contains the methods of estimating the parameters of the different models. The results of the study are found in Section 5. Corcoran and Read [3] conducted a Monte Carlo study to compare several reliability growth models that includes two of the models considered in this paper. However, there are several differences in the methodology used in the two studies.

## 2. Description of the Reliability Growth Models.

Model I: Duane [4] developed this model based on his analysis of large quantities of data on various types of airborne equipment. If  $T$  is the accumulated test time on a given piece of equipment, and  $\lambda_T$  is the instantaneous failure rate at time  $T$  then, Duane's model states that

$$\lambda_T = \beta(1-\alpha)T^{-\alpha}$$

where  $\beta$  and  $\alpha$  are unknown parameters.

Model II: Suppose that (i) items are put on test sequentially and after each failure an appropriate corrective action is taken and (ii) the time to failure on the  $i^{\text{th}}$  test has an exponential distribution with parameter  $\lambda_i$ . Then, Weiss [5] assumes that the parameters  $\lambda_i$  satisfy the equation

$$\lambda_i = \frac{B + i}{Ai}$$

for some unknown parameters  $A$  and  $B$ .

Model III: Let

$p_r = P["\text{mission success} \text{ after the } r^{\text{th}} \text{ failure has been observed and appropriate corrective action has been taken}]$

Chernoff and Woods [ 2 ] model  $p_r$  as an increasing function of  $r$  defined by

$$p_{r+1} = 1 - e^{-(\alpha + \beta r)}$$

Model IV: Wolman [ 6 ] developed this model assuming that

- (i) testing is conducted sequentially and each trial results in a "mission success" or a "mission failure"
- (ii) on each trial one of two types of failures can occur; the first type is an inherent or non-correctable failure and the second type of failure is a transient failure that can be eliminated by corrective action and
- (iii) whenever a test results in a "failure" an attempt is made to eliminate the failure mode and the attempt may result in a) either a complete elimination b) a partial elimination or c) no effect on the failure mode.

Let

$r$  = probability of an inherent failure, constant from test to test

$q$  = probability of a transient failure at the beginning of the test program

$\beta_i$  = factor by which the corrective action after the  $i^{\text{th}}$  failure, reduces the probability of a transient failure  $i-1$

$\alpha_i = \prod_{j=0}^i \beta_j$  the cumulative factor by which the initial probability of a transient failure is reduced as a result of the corrective actions after each of the  $i$  failures and

$\beta_{n+1}(i)$  = probability of a "mission success" on the  
( $n+1$ )st test if  $i$  failures are observed  
in the first  $n$  tests

The Wolman model assumes that

$$p_{n+1}(i) = 1 - r - q\alpha_i$$

Wolman points out that in many situations a good approximation is obtained by assuming that the improvement factors  $\beta_i$  are all equal to a constant value  $\beta$ . In this case, the model equation reduces to

$$p_{n+1}(i) = 1 - r - q\beta^i$$

The latter version with  $\beta = .2, .5$  and  $.8$  is used in this study.

### 3. Monte Carlo Procedure

In order to compare the growth models described in the previous section, five decreasing sequences of  $\lambda$  (failure rate) values are selected. Each sequence consists of 16 values ranging between .70 and .05; this represents a growth in reliability, assuming that all times are measured in mission units, from an initial value of .50 to a final value of .95 approximately. The growth patterns represented in these sequences vary from a very rapid growth to a very slow growth.

For each of the 16  $\lambda$ 's in a given sequence 10 exponential failure times  $t_{ij}$   $i=1,2,\dots,16$ ;  $j=1,2,\dots,10$  are generated. Since a reliability growth model is

to predict the  $\lambda$  values that are in practical situations unknown, the generated data is used to first estimate the  $\lambda$  values in a given sequence in two ways. In the first method the  $t_{ij}$ 's are treated as the observed exponential failure times and in the second case the  $t_{ij}$ 's are converted into bernoulli data by noting whether each  $t_{ij}$  exceeds 1 or not. This is done to study the effect of different data types on the prediction capabilities of the growth models. The estimated  $\lambda$  values are used to empirically determine the unknown parameters of the growth models using least squares techniques. To account for statistical variability the whole process described above is replicated 100 times resulting in 100 empirical growth curves for each of the  $\lambda$  sequences and each of the reliability growth models.

If  $\lambda_i$ ,  $i = 1, 2, \dots, 16$  are the failure rates in a given sequence 100 predicted values  $\bar{\lambda}_i^*$  are obtained from the empirical curves for each of the models. The "goodness" of a reliability growth model is now measured using the following three measures:

$$M_1 = \sum_{i=1}^{16} \left| \frac{\lambda_i - \bar{\lambda}_i^*}{\lambda_i} \right|$$

$$M_2 = \sum_{i=1}^{16} \left[ \frac{\lambda_i - \bar{\lambda}_i^*}{\lambda_i} \right]^2$$

$$M_3 = \max_i \left| \frac{\lambda_i - \bar{\lambda}_i^*}{\lambda} \right|$$

where  $\bar{\lambda}_i^*$  is the average of the 100 predicted values  $\bar{\lambda}_i^*$ . All three measures are relative measures i.e., each difference  $\lambda_i - \bar{\lambda}_i^*$  is divided

by the true value  $\lambda_i$  in computing the measures. The reason for choosing relative measures is that it would be important for a growth model to predict the small  $\lambda$  values at the end of a sequence rather well. In other words, these measures attach more weight to the differences  $\lambda_i - \bar{\lambda}_i^*$  at the end of each decreasing sequence.

#### 4. Estimation of Parameters

As indicated earlier the generated failure times  $t_{ij}$ ,  $j = 1, 2, \dots, 10$  are used to estimate the failure rates  $\lambda_i$  in two different ways. For the case where the  $t_{ij}$ 's are used directly the estimator of  $\lambda_i$  is the maximum likelihood estimator  $\hat{\lambda}_i = 1/\bar{t}_i$  where  $\bar{t}_i = \frac{1}{10} \sum_{j=1}^{10} t_{ij}$ . Since models III and IV describe the growth in  $p_i$  the probability of a mission success,  $\hat{\lambda}_i$  is converted into an estimator of  $p_i$  by noting that

$$\begin{aligned} p_i &= P[\text{mission success on } i^{\text{th}} \text{ trial}] \\ &= P[t_{ij} \geq 1] = e^{-\lambda_i} \end{aligned}$$

Thus, the maximum likelihood estimator of  $p_i$  is  $\hat{p}_i = e^{-\hat{\lambda}_i}$ . In the second case, the  $t_{ij}$ 's are transformed to bernoulli data and the  $\lambda_i$  and  $p_i$  are estimated from the transformed data as follows:

Let

$$x_{ij} = \begin{cases} 1 & \text{if } t_{ij} \geq 1 \\ 0 & \text{if } t_{ij} < 1 \end{cases}$$

and let  $Y_i = \sum_{j=1}^k X_{ij}$ . Then,  $Y_i$  has a binomial distribution with parameters  $p_i$  and  $k$ . The maximum likelihood estimators of  $p_i$  and  $\lambda_i$  based on  $Y_i$  are  $\hat{p}_i = \frac{Y_i}{k}$  and  $\hat{\lambda}_i = -\ln \hat{p}_i$  respectively. The methods for obtaining the empirical growth curves and the predicted  $\lambda$  values from the growth curves are now described. Since these methods are the same for both sets of estimators  $\hat{\lambda}_i$ ,  $\hat{p}_i$  and  $\tilde{\lambda}_i$ ,  $\tilde{p}_i$  the symbols  $\tilde{\lambda}_i$ ,  $\tilde{p}_i$  will be used to denote either of the two sets of estimators. Also, the symbols  $\lambda_i^*$ ,  $p_i^*$  denote the  $\lambda_i$ ,  $p_i$  values as predicted by the empirical growth curves.

For model I the equation describing the relationship between the failure rate  $\lambda_T$  and the total test time  $T$  is

$$\lambda_T = \beta(1-\alpha)T^{-\alpha}$$

or equivalently

$$\ln \lambda_T = \ln[\beta(1-\alpha)] - \alpha \ln T$$

The parameters  $\alpha$  and  $\beta$  are estimated by replacing  $\lambda_T$  with  $\tilde{\lambda}_i$  and  $T$  with  $T_i = \sum_{\ell=1}^{10} \sum_{j=1}^{t_{\ell i}} t_{\ell j}$  in the latter equation and then applying standard regression methods. Once the least squares estimators  $\hat{\alpha}$  and  $\hat{\beta}$  are obtained, the predicted values  $\lambda_i^*$  of the failure rates  $\lambda_i$  are given by

$$\lambda_i^* = \hat{\beta}(1 - \hat{\alpha})T_i^{-\hat{\alpha}} \quad i = 1, 2, \dots, 16$$

The parameters  $A$  and  $B$  of model II are estimated using regression methods on the relationship

$$\tilde{\lambda}_i = \frac{B}{A} \frac{1}{i} + \frac{1}{A} \quad i = 1, 2, \dots, 16$$

and the predicted values are

$$\lambda_i^* = \frac{\hat{B} + i}{\hat{A}i} \quad i = 1, 2, \dots, 16$$

For model III the regression equation has the form

$$-\ln(1 - \tilde{p}_i) = \alpha + \beta(i-1) \quad i = 1, 2, 3, \dots, 16$$

and for model IV the equation is

$$1 - p_i = r + q\beta^{i-1} \quad i = 1, 2, \dots, 16 \quad \text{with } \beta \text{ known .}$$

The rest of the procedure is the same as for models I and II.

## 5. SIMULATION RESULTS

The results of the simulation are presented in tables 1.1 to 1.5.

The five tables correspond to the five sets of 16  $\lambda$  values used in generating the failure times. These  $\lambda$  values are listed at the top of the tables. As indicated earlier the five sets chosen represent reliability growth ranging between a slow growth to a rapid growth. Also, the selection of the  $\lambda$ 's in these sets are made so as to generally follow the growth indicated by the models that are being tested. The entries in the tables are the observed values of the three measures  $M_1, M_2, M_3$  for each of the four models with three cases under model IV.

It may be observed from tables 1.1 to 1.5 that with time to failure data model I performed better than the others in all cases except one (table 1.4) where model III was best. For attributes data model IV with  $\beta = .8$  appears to outperform the other models. This result coincides with a similar conclusion by Bresenham [1].

In order to study the case where exactly one item is put on test at any time, the simulation was rerun generating only one failure time  $t_{ij}$  instead of 10 failure times  $t_{ij}$  for each of the 16  $\lambda$ 's in a given sequence. There was a difference in the generation of attribute data

from the failure times. Let  $X_i$  denote the largest integer less than or equal to  $t_i$ , the time to failure on the  $i^{\text{th}}$  trial with failure rate  $\lambda_i$ ,  $i = 1, 2, \dots, 16$ . Since the  $t_i$  are assumed to be measured in mission units  $X_i$  is equivalent to the number of full missions completed before the occurrence of a failure. Thus,  $X_i$  has a geometric distribution with parameter  $p_i$ , the probability of a mission success on the  $i^{\text{th}}$  trial. The maximum likelihood estimator of  $p_i$  is  $\hat{p}_i = 1/(X_i + 1)$ . Except for this difference in the way the  $p_i$ 's are estimated the rest of the procedure is the same as before. The results of this second study are in tables 2.1 to 2.5. In this case, model III appears to do uniformly well with time to failure data and with attributes data model III and model IV with  $\beta = .8$  seems to perform better than the others.

On the basis of this study, the following general conclusions may be drawn:

- (1) with time to failure data and several items on test simultaneously, model I appears to predict reliability growth accurately
- (2) with attributes data and several items on test model IV with  $\beta = .8$  performs well
- (3) with time to failure data and one item on test model III is a good model to use and
- (4) with attributes data and one item on test either model III or model IV with  $\beta = .8$  may be used.

Graphs comparing true growth curves (curves down through the  $\lambda$  values used in the study) with the curves estimated by model I using time data and model IV using attribute data with  $\beta = .8$  are presented in Figures 1.1 to 1.5. Figures 2.1 to 2.5 compare the true growth curves with the curves estimated by model III using time and attribute data and model IV using attribute data with  $\beta = .8$ .

The computer programs used in the study are attached.

TABLE 1.1

True Values of Lambda: .702 .434 .320 .255 .213 .183 .161  
 .144 .130 .118 .109 .101 .0936 .0876 .0823 .0776

MODEL NUMBER	TIME DATA			ATTRIBUTE DATA		
	M1	M2	M3	M1	M2	M3
Model I	0.69	0.04	0.10	9.46	6.23	0.99
Model II	2.45	0.65	0.40	2.28	0.54	0.38
Model III	2.27	0.44	0.37	7.91	4.66	.778
Model IV $\beta = (.2)$	7.49	5.16	1.13	6.70	4.04	0.99
Model IV $\beta = (.5)$	4.52	2.13	0.79	4.09	1.57	0.67
Model IV $\beta = (.8)$	1.50	0.24	0.23	1.55	0.18	0.16

TABLE 1.2

True Values of Lambda: .700 .573 .508 .466 .435 .411 .392  
 .376 .362 .350 .340 .330 .322 .314 .307 .301

MODEL NUMBER	TIME DATA			ATTRIBUTE DATA		
	M1	M2	M3	M1	M2	M3
Model I	0.68	0.04	0.11	2.69	0.50	0.22
Model II	1.67	0.27	0.23	1.35	0.18	0.20
Model III	1.04	0.09	0.15	2.76	0.58	0.32
Model IV $\beta = (.2)$	2.67	0.58	0.33	2.21	0.40	0.25
Model IV $\beta = (.5)$	1.85	0.32	0.27	1.58	0.20	0.18
Model IV $\beta = (.8)$	1.09	0.08	0.11	0.39	0.01	0.06

TABLE 1.3

True Values of Lambda: .700 .353 .238 .180 .145 .122 .106  
 .0933 .0837 .0760 .0697 .0644 .0600 .0562 .0529 .0500

MODEL NUMBER	TIME DATA			ATTRIBUTE DATA		
	M1	M2	M3	M1	M2	M3
Model I	0.78	0.06	0.11	11.00	8.04	0.82
Model II	1.61	0.16	0.11	0.92	0.06	0.08
Model III	3.17	0.86	0.50	10.20	6.94	0.86
Model IV $\beta = (.2)$	8.42	6.57	1.26	7.33	4.88	1.08
Model IV $\beta = (.5)$	4.22	1.93	0.76	3.59	1.28	0.61
Model IV $\beta = (.8)$	3.87	1.16	0.43	3.81	1.13	0.44

TABLE 1.4

True Values of Lambda: .700 .564 .460 .379 .315 .263 .221  
 .186 .157 .132 .112 .0953 .0810 .0689 .0587 .0500

MODEL NUMBER	TIME DATA			ATTRIBUTE DATA		
	M1	M2	M3	M1	M2	M3
Model I	1.88	0.48	0.55	11.60	13.80	2.91
Model II	8.32	8.93	1.85	8.01	8.16	1.78
Model III	0.56	0.02	0.06	11.90	19.80	3.85
Model IV $\beta = (.2)$	14.10	24.30	3.04	12.90	20.10	2.77
Model IV $\beta = (.5)$	10.60	14.20	2.39	9.75	11.40	2.15
Model IV $\beta = (.8)$	2.45	0.92	0.65	1.95	0.47	0.48

TABLE 1.5

True Values of Lambda: .700 .456 .315 .227 .170 .132 .106  
 .0880 .0757 .0672 .0613 .0572 .0544 .0524 .0510 .0500

MODEL NUMBER	TIME DATA			ATTRIBUTE DATA		
	M1	M2	M3	M1	M2	M3
Model I	1.23	0.13	0.14	11.80	9.74	1.54
Model II	3.67	1.07	0.38	3.73	1.11	0.39
Model III	3.70	1.09	0.41	9.88	6.60	0.88
Model IV $\beta = (.2)$	11.80	12.10	1.47	10.80	10.20	1.34
Model IV $\beta = (.5)$	6.57	3.93	0.88	6.01	3.11	0.78
Model IV $\beta = (.8)$	3.68	1.12	0.47	3.47	1.07	0.53

TABLE 2.1

True Values of Lambda: .702 .434 .320 .255 .213 .183 .161  
 .144 .130 .118 .109 .101 .0936 .0876 .0823 .0776

MODEL NUMBER	TIME DATA			ATTRIBUTE DATA		
	M1	M2	M3	M1	M2	M3
Model I	19.20	43.10	5.44	5.71	6.15	2.21
Model II	13.20	11.30	10.80	30.7	83.8	4.83
Model III	9.63	6.37	0.76	4.56	1.68	.559
Model IV $\beta = (.2)$	40.50	297.00	15.40	16.50	85.00	6.26
Model IV $\beta = (.5)$	36.70	229.00	13.50	13.50	63.30	7.74
Model IV $\beta = (.8)$	23.00	35.00	35.00	3.88	1.37	0.57

TABLE 2.2

True Values of Lambda: .700 .573 .508 .466 .435 .411 .392  
 .376 .362 .350 .340 .330 .322 .314 .307 .301

MODEL NUMBER	TIME DATA			ATTRIBUTE DATA		
	M1	M2	M3	M1	M2	M3
Model I	18.60	38.40	5.02	3.73	3.51	1.78
Model II	118.00	877.00	8.35	8.68	14.10	3.15
Model III	6.36	2.70	0.49	2.16	.354	.240
Model IV $\beta = (.2)$	24.60	169.00	12.60	7.80	24.40	4.86
Model IV $\beta = (.5)$	24.40	164.00	12.40	7.27	23.10	4.75
Model IV $\beta = (.8)$	13.80	12.70	1.61	3.83	3.52	1.76

TABLE 2.3

True Values of Lambda: .700 .353 .238 .180 .145 .122 .106  
 .0933 .0837 .0760 .0697 .0644 .0600 .0562 .0529 .0500

MODEL NUMBER	TIME DATA			ATTRIBUTE DATA		
	M1	M2	M3	M1	M2	M3
Model I	19.20	41.20	5.20	5.47	6.17	2.26
Model II	127.00	1010.00	8.68	50.80	216.00	6.40
Model III	11.10	8.53	0.93	4.79	1.76	0.64
Model IV $\beta = (.2)$	45.80	303.00	14.60	15.90	60.50	7.30
Model IV $\beta = (.5)$	41.40	247.00	13.40	10.40	29.50	5.12
Model IV $\beta = (.8)$	25.90	44.50	2.12	3.54	.834	0.29

TABLE 2.4

True Values of Lambda: .700 .564 .460 .379 .315 .263 .221  
 .186 .157 .132 .112 .0953 .0810 .0689 .0587 .0500

MODEL	TIME DATA			ATTRIBUTE DATA		
	M1	M2	M3	M1	M2	M3
Model I	21.00	70.00	7.49	8.13	11.90	2.91
Model II	182.00	2670.00	25.50	27.00	77.50	5.02
Model III	9.91	7.21	12.90	4.67	2.46	0.93
Model IV $\beta = (.2)$	49.20	433.00	17.40	25.90	167.00	11.90
Model IV $\beta = (.5)$	47.50	475.00	19.30	23.30	157.00	11.80
Model IV $\beta = (.8)$	28.40	81.60	5.72	8.54	10.10	1.91

TABLE 2.5

True Values of Lambda: .700 .456 .315 .227 .170 .132 .106  
 .0880 .0757 .0672 .0613 .0572 .0544 .0524 .0510 .0500

MODEL NUMBER	TIME DATA			ATTRIBUTE DATA		
	M1	M2	M3	M1	M2	M3
Model I	19.50	45.10	5.62	6.08	6.46	2.27
Model II	145.00	1370.00	11.10	48.50	205.00	6.34
Model III	10.90	9.22	1.11	5.50	2.21	0.59
Model IV $\beta = (.2)$	52.20	421.00	17.40	25.10	204.00	13.90
Model IV $\beta = (.5)$	46.30	353.00	16.50	18.20	124.00	10.90
Model IV $\beta = (.8)$	25.90	45.70	2.06	3.42	0.90	0.35

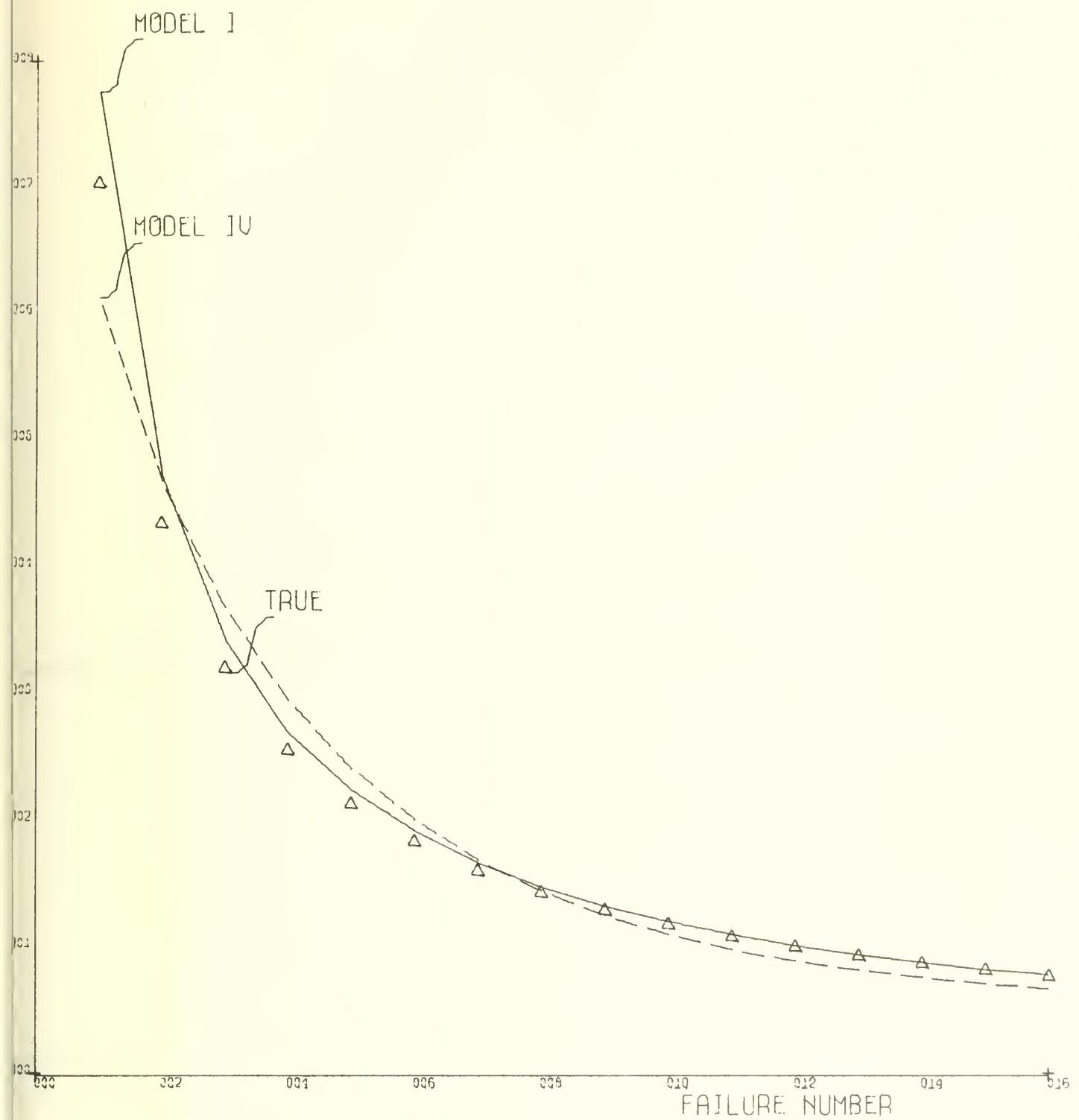


FIGURE 1.1 GRAPH OF PREDICTED FAILURE RATES FOR MODEL I, MODEL IV(.8), AND TRUE RATE

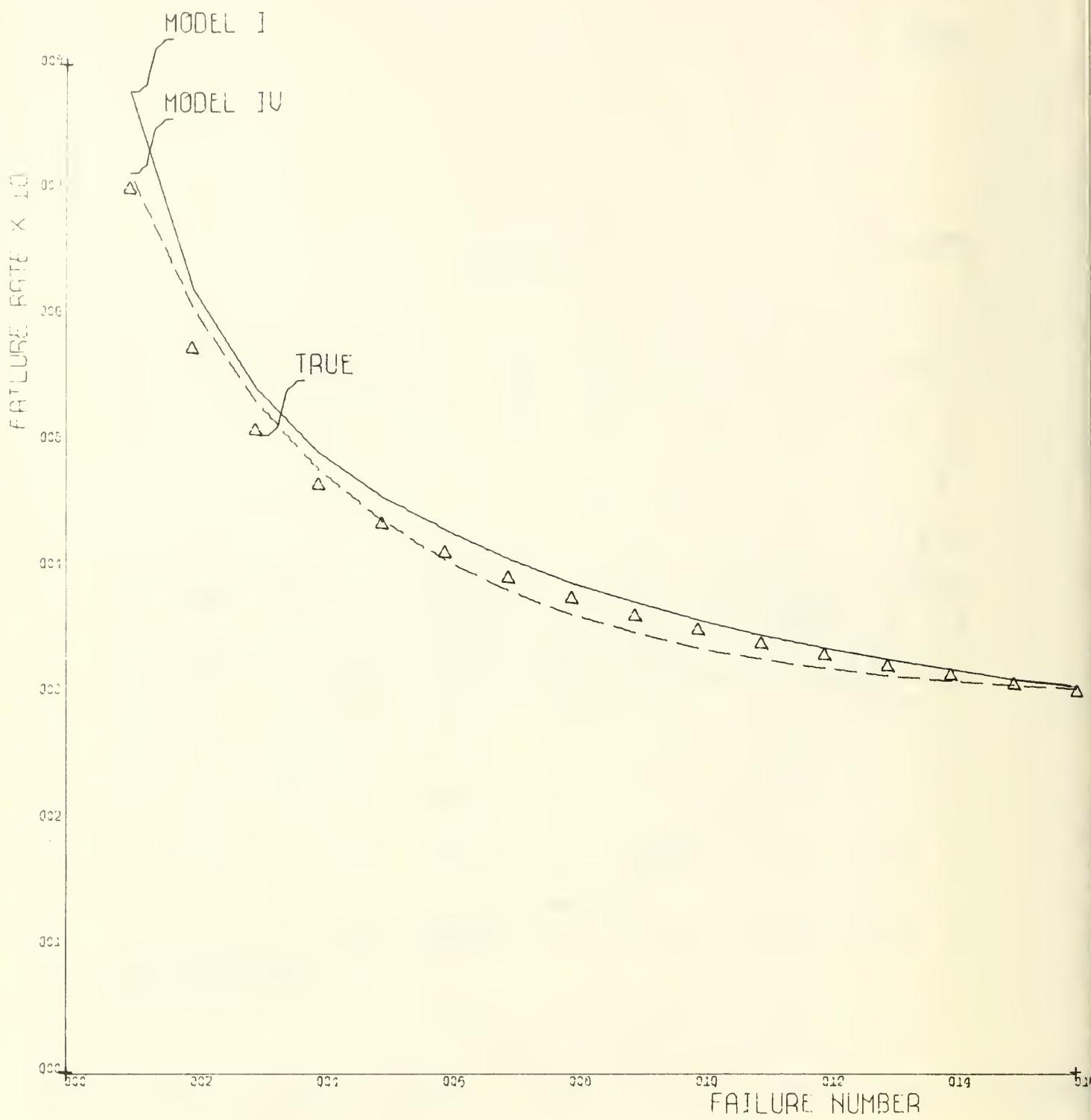


FIGURE 1.2 GRAPH OF PREDICTED FAILURE RATES FOR MODEL I, MODEL IV(.8), AND TRUE RATE

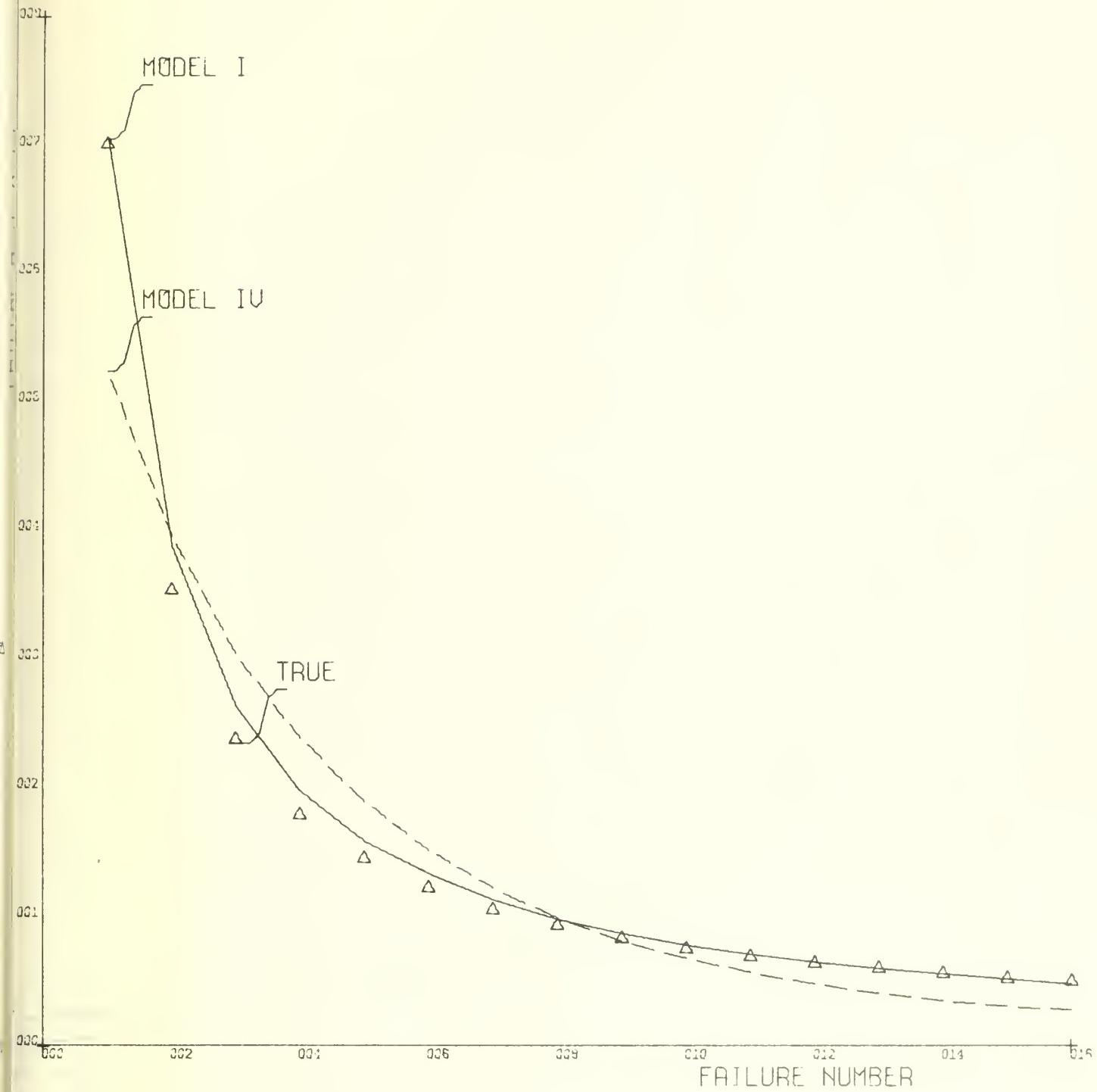


FIGURE 1.3 GRAPH OF PREDICTED FAILURE RATES FOR MODEL I, MODEL IV(.8), AND TRUE RATE

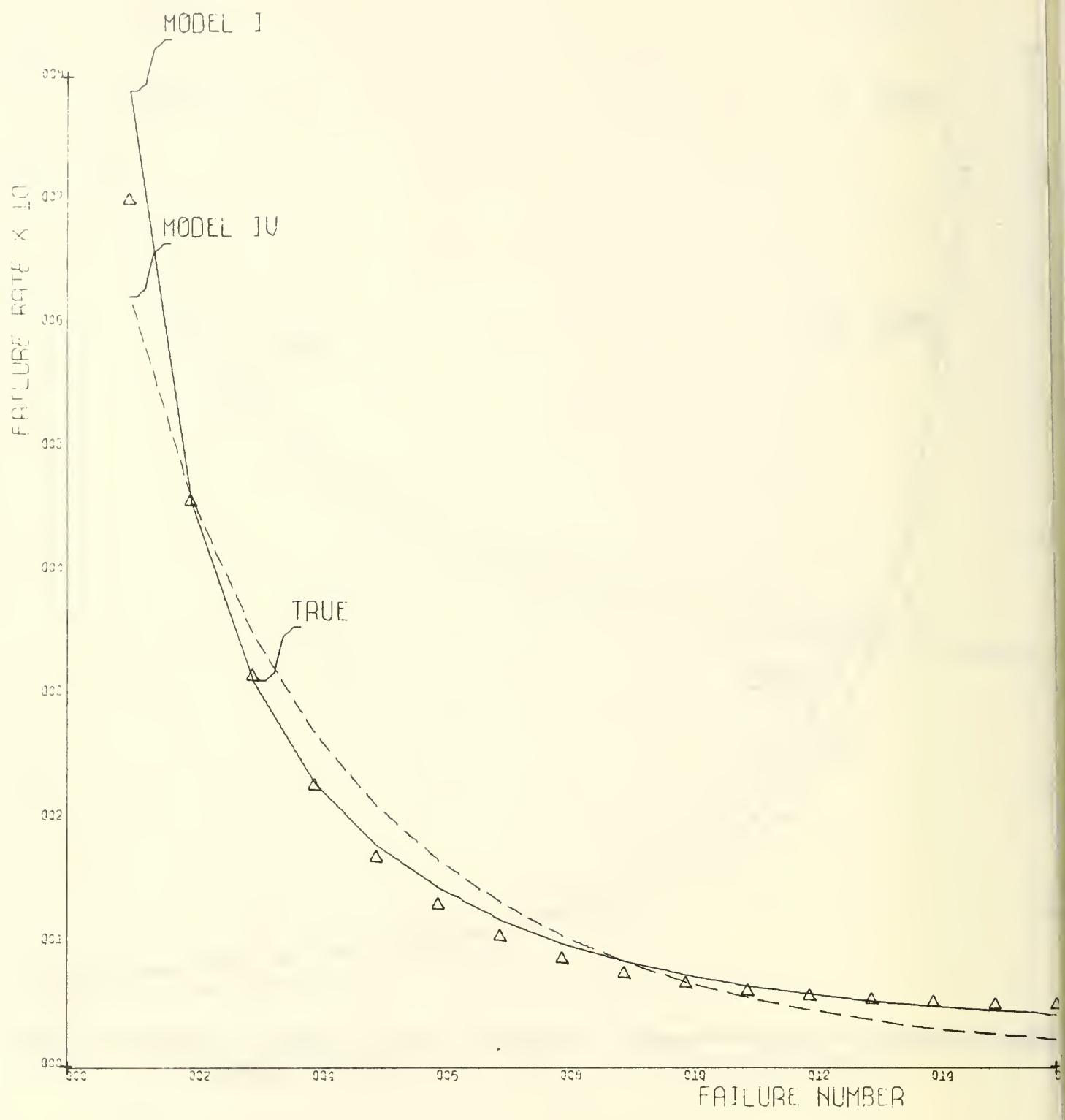


FIGURE 1.4 GRAPH OF PREDICTED FAILURE RATES FOR MODEL I, MODEL IV(.8), AND TRUE RA

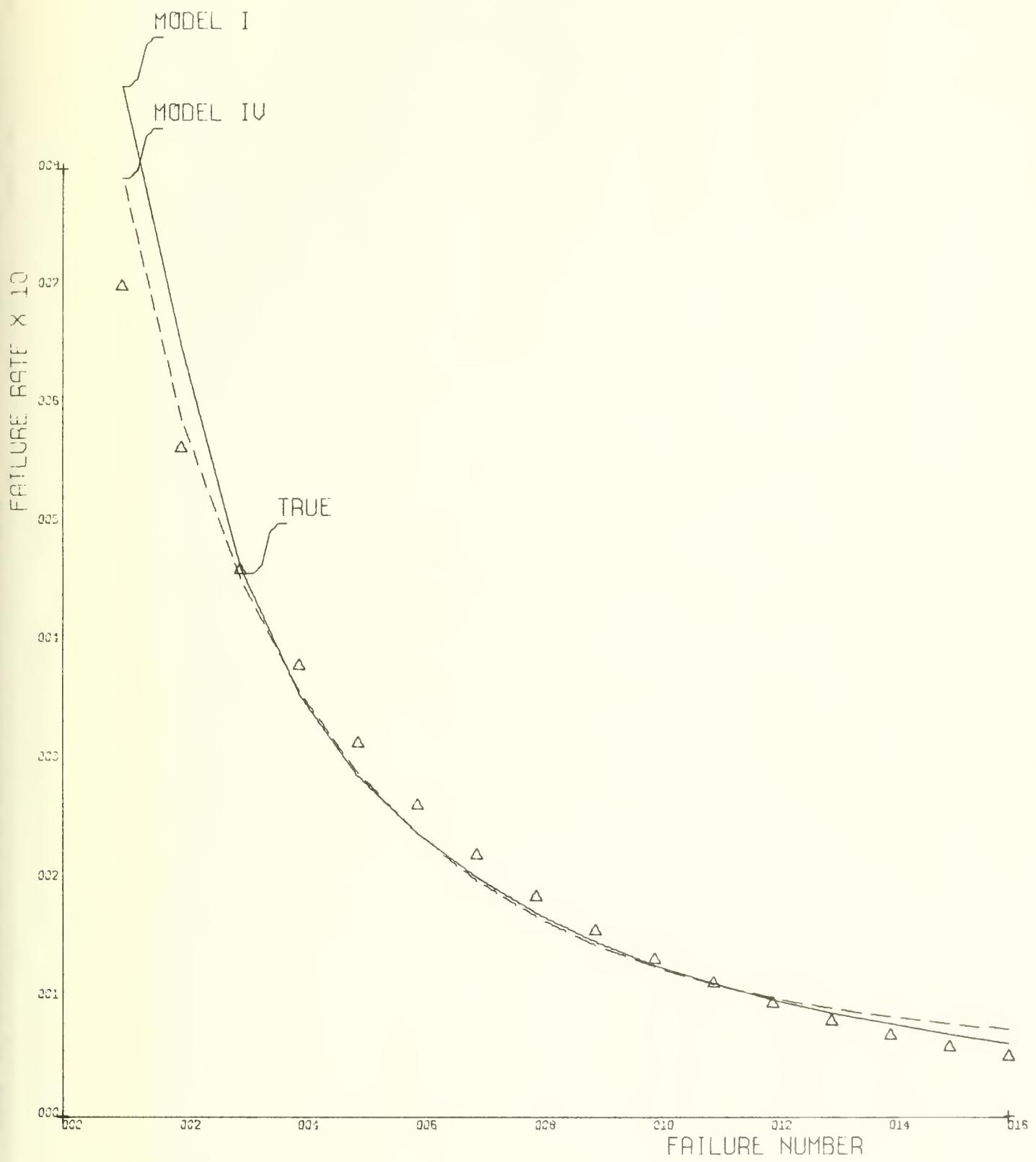


FIGURE 1.5 GRAPH OF PREDICTED FAILURE RATES FOR MODEL I, MODEL IV(.8), AND TRUE RATE

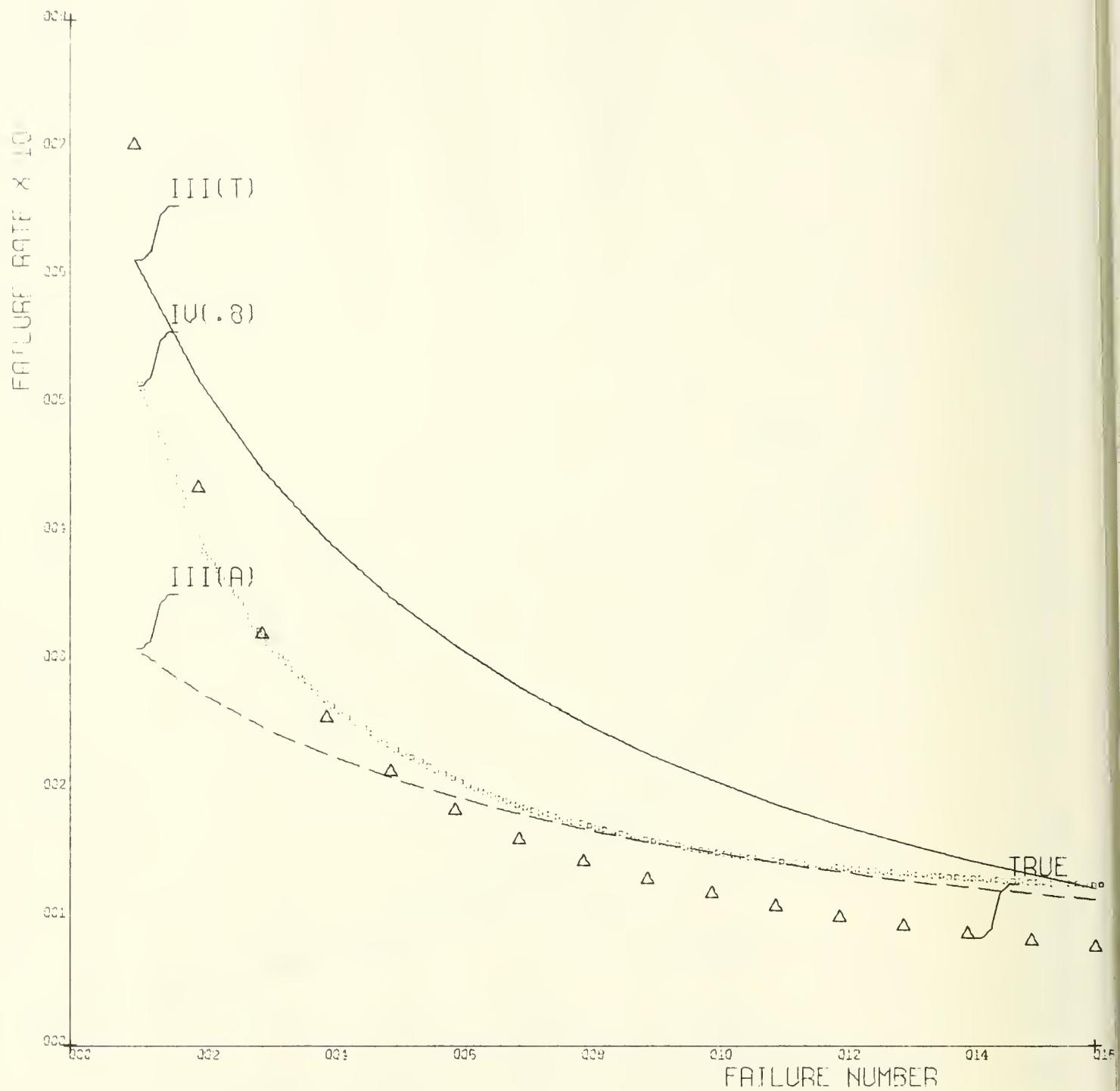


FIGURE 2.1 GRAPH OF FAILURE RATES PREDICTED BY MODELS III(T&A), IV(.8) AND TRUE RF

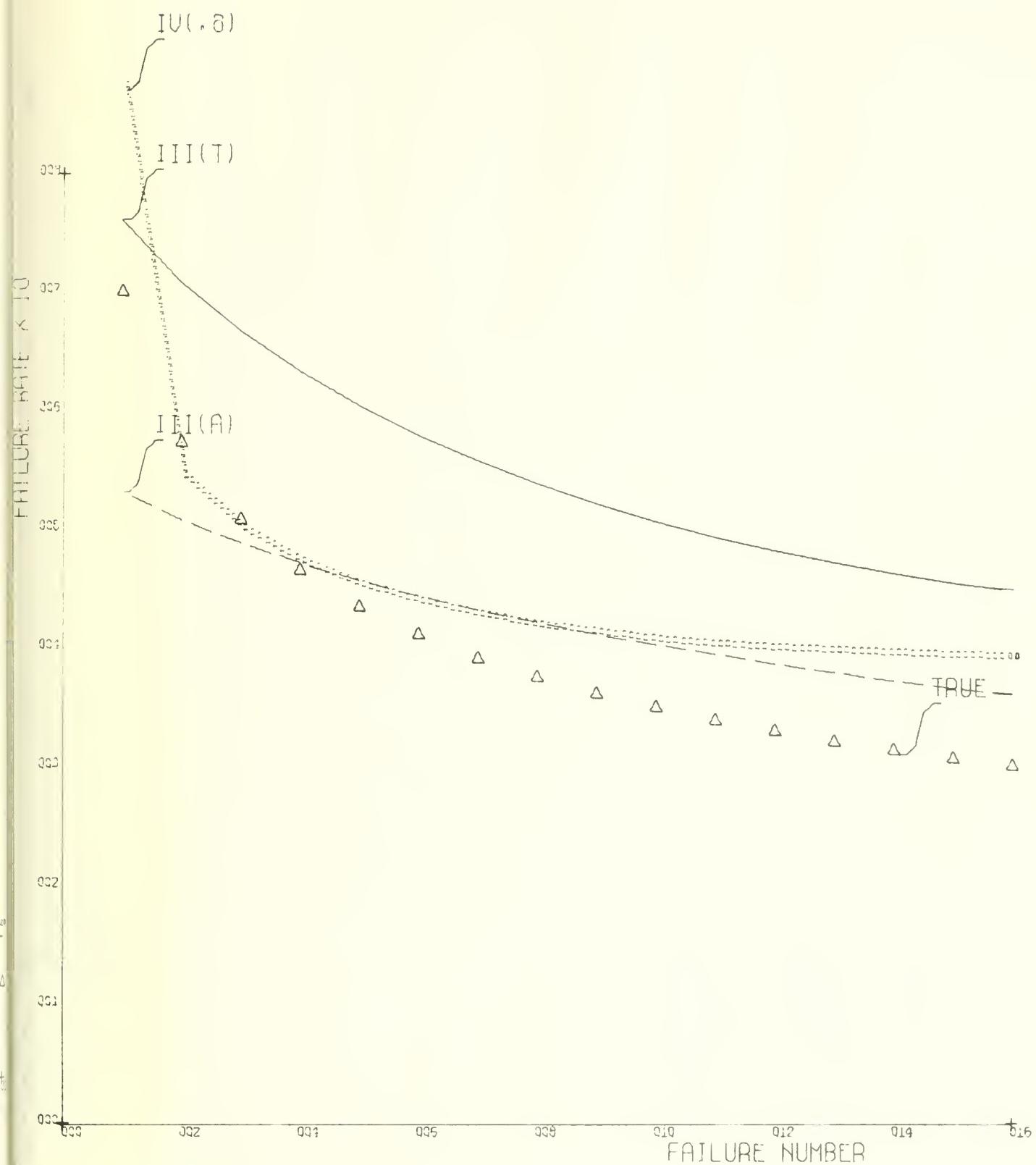


FIGURE 2.2 GRAPH OF FAILURE RATES PREDICTED BY MODELS III(T&A), IV(.8) AND TRUE RATE

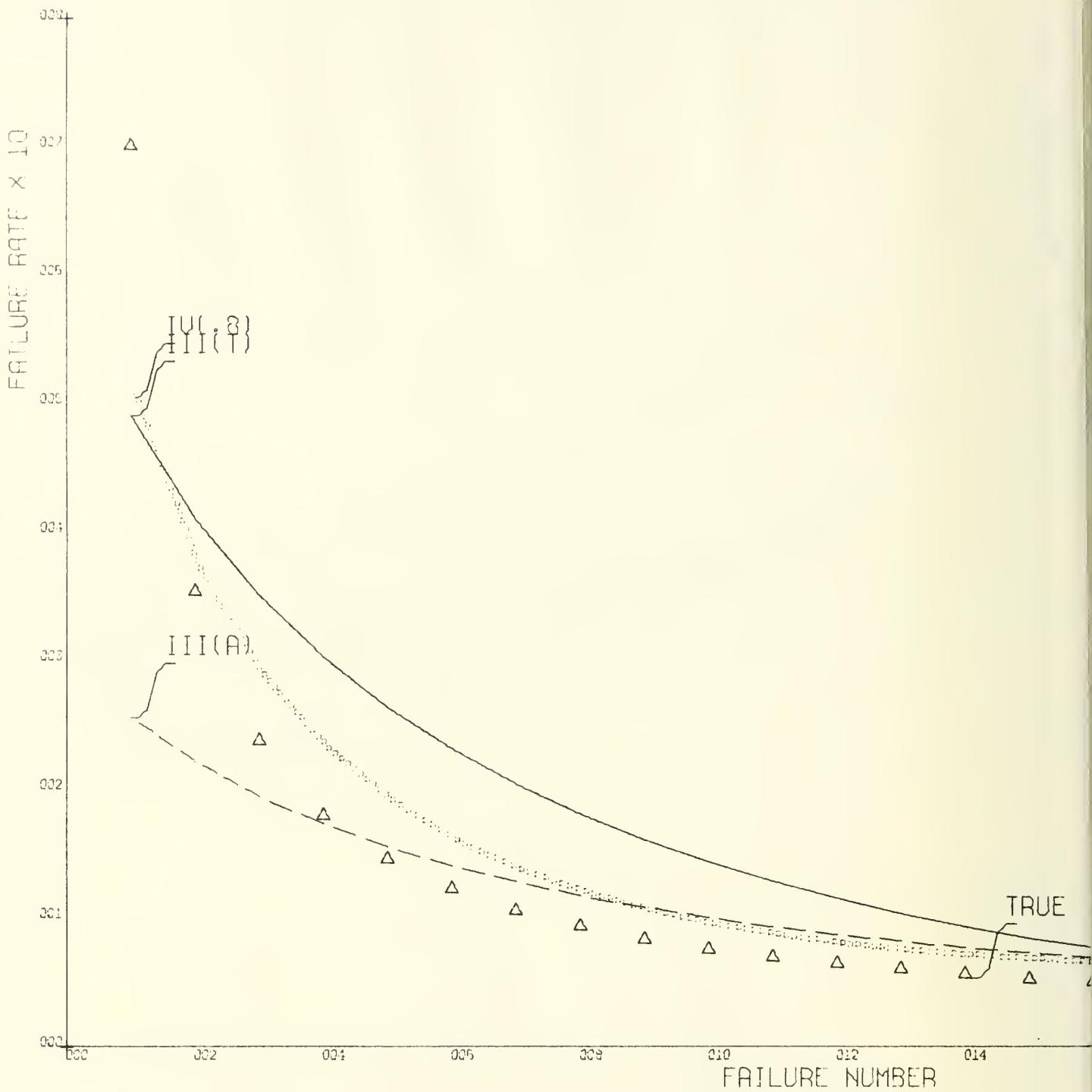


FIGURE 2.3 GRAPH OF FAILURE RATES PREDICTED BY MODELS III(T&A), IV(.8) AND TRUE F

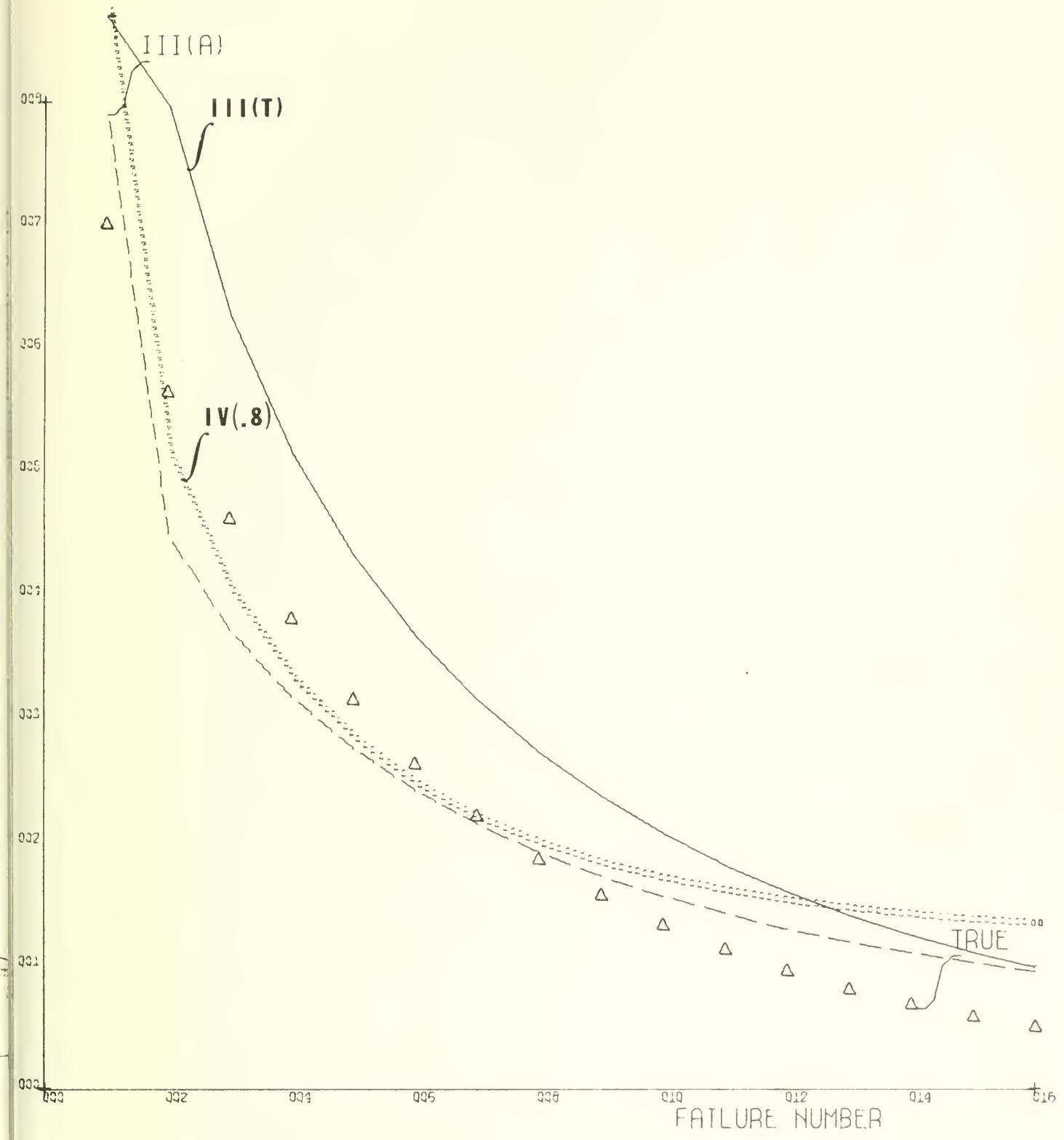


FIGURE 2.4 GRAPH OF FAILURE RATES PREDICTED BY MODELS III(T&A), IV(.8) AND TRUE RATE

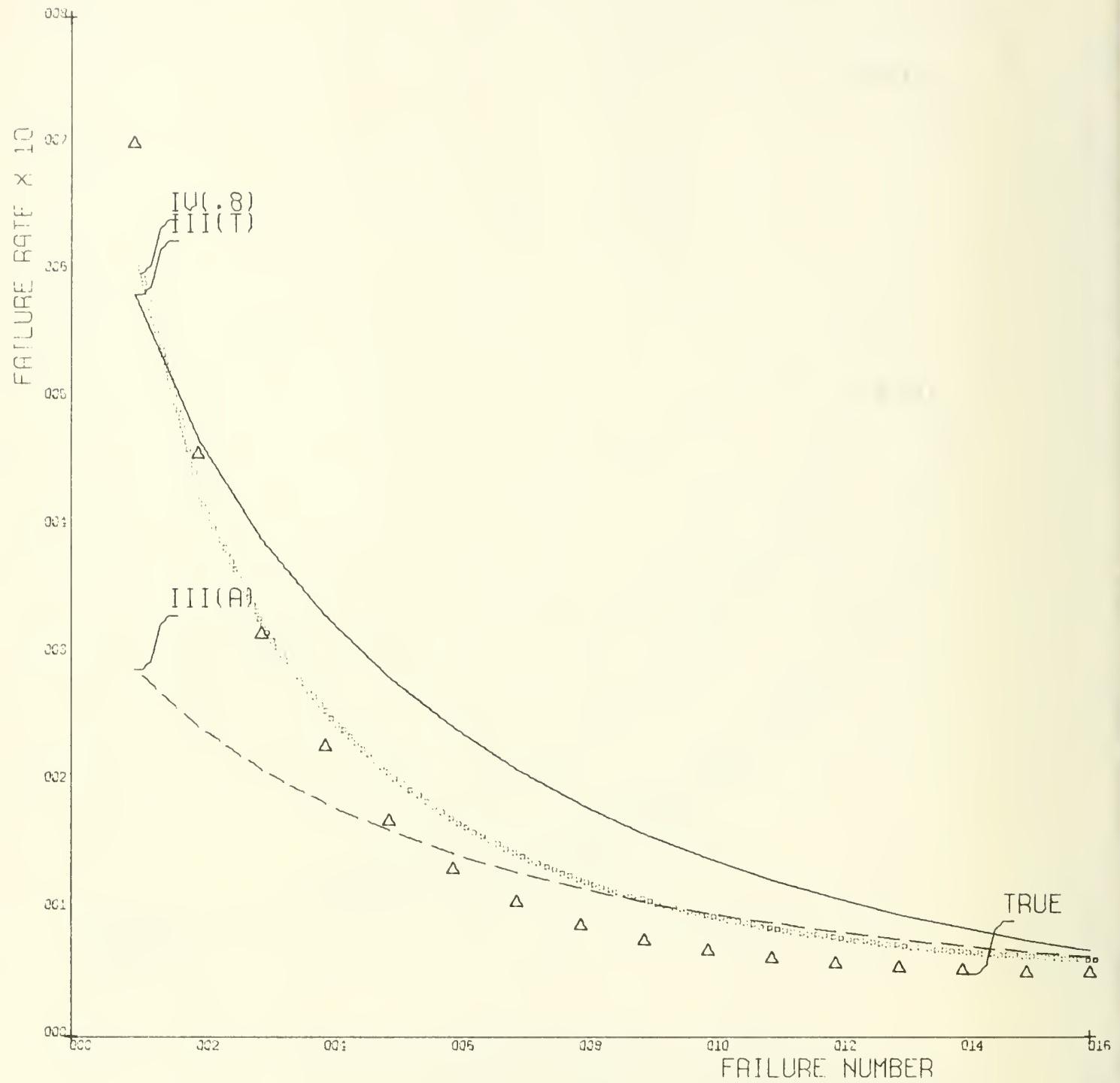


FIGURE 2.5 GRAPH OF FAILURE RATES PREDICTED BY MODELS III(T&A), IV(.8) AND TRUE RATE

### References

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- [2] Chernoff, Herman and Woods, W. Max, "Reliability Growth Models -- Analysis and Applications," CEIR, Inc., file memo dated February 26, 1962.
- [3] Corcoran, W. J. and Read, R. R., "Comparison of Some Reliability Growth Estimation and Prediction Schemes," UTC 2140-ITR Addendum, United Technology Center, Sunnyvale, California, June 1967.
- [4] Duane, J. T., "Learning Curve Approach to Reliability Monitoring," IEEE Transactions on Aerospace, Vol. 2, No. 2, 1964, pp. 563-566.
- [5] Weiss, H. K., "Estimation of Reliability Growth in A Complex System with a Poisson-type Failure," Operations Research, Vol. 4, 1956, pp. 532-545.
- [6] Wolman, W., "Problems in System Reliability Analysis,: Statistical Theory of Reliability, University of Wisconsin, 1963.

```

// EXEC FORTCLG,REGION=200K
//FORT.SYSIN DD *

C RELIABILITY GROWTH SIMULATION PROGRAM
C
***** SAMPLE DATA DECK FOLLOWS THIS PROGRAM
***** DATA IS INPUT VIA NAMELISTS LAMDA AND TIMES
***** &LAMDA NUM=NN,NOFMT=FF &END
***** WHERE NN IS THE NUMBER OF LAMDAS
***** AND FF IS THE NUMBER OF FORMAT CARDS
***** WHICH FOLLOW AS ("FORMAT ")
***** TRUE LAMDA VALUES FOLLOW FORMAT CARDS
***** &TIMES NOREP=RR, ISEED=SSSSSSS &END
***** WHERE RR = NUMBER OF REPLICATIONS AND
***** SSSSSS = THE SEED FOR THE SIMULATION

( THE NEXT 4 KEY WORDS CAUSE THE MODELS
  INDICATED TO PREDICT THE TRUE LAMDAS)
GENERAL=>MODEL_I
WEISS=>MODEL_II
WOODS=>MODEL_III
WOLMAN=>MODEL_IV COL 41 - 50 ARE FOR BETA

RESET => RESETS ALL PRINT OPTIONS TO INITIAL CONDITIONS
COL 41 - 50 = FIRST MEASURE PRINTED
COL 51 - 60 = NUMBER OF SIMULATIONS
COL 71 - 80 = OUTPUT DEVICE FOR MEASURES
COL 61 - 70 = LAST MEASURE PRINTED
PRINT => PRINT OUT RESULTS TO DATE
A CARD WITH NO KEY WORD CAUSES THE PREDICTIONS
BASED ON THE RANDOM NUMBERS GENERATED
TO BE USED AS IF THEY WERE THE RESULT OF
ONE OF THE MODELS

```

OUTPUT DEVICE 6 "PRINTER"  
 RESULTS OF EACH SIMULATION  
 USUALLY DUMMIED FOR LARGE NUMBER OF SIMULATIONS  
 DEVICE IN COL 71 - 80 OF RESET CARD  
 FINAL RESULTS FOR MEASURES SPECIFIED  
 DEVICE 9 "CARDS"  
 PREDICTED LAMDAS ARE OUTPUT FOR PLOTTING

THIS PROGRAM CALCULATES AND OUTPUTS OTHER QUANTITIES THAN THE ONES MENTIONED ABOVE FOR FUTURE RESEARCH.

PROGRAMS CALLED OTHER THAN THOSE PRESENTED BELOW ARE IN THE NPS PROGRAM LIBRARY OF SPECIAL SIGNIFICANCE IS THE PROGRAM "LRANDOM" WHICH GENERATES OUR RANDOM NUMBERS

```

REAL*8 HOLDX(300)
REAL*8 HOLDX(300)
DIMENSION FITTED(301)
DIMENSION AMDA(301)
REAL*4 MODEL(10)
DIMENSION PARM(4)
REAL*4 MEAS(8,6)
REAL*4 WOLM,'WOLM'/
REAL*4 NEWL/'NEW'/
REAL*4 WEIS/'WEIS'/
REAL*4 MOOR/'MOOR'/
REAL*4 GENE/'GENE'/
REAL*4 FINI/'FINI'/
REAL*4 WOOD/'WOOD'/
REAL*4 RESE/'RESE'/'PRIN'/'PRIN'/
LOGICAL NOUT=6
CONTINUE NSIM=1
NMEAS=3

```

```

IF(NMEND .LE. 0) NMMEAS=3
IF(NSIM .LE. 0) NSEND=NMEAS
IF(NSIM .LE. 0) NSIM=1
IF(NOUT .LE. 0) NOUT=6
READ(5,100) END=999 MODEL, PARM
FORMAT(5,100) END=999 MODEL, PARM
100 WRITE(6,101) MODEL
FORMAT(6,101) MODEL
101 IF(MODEL(1) .NE. RESE) GO TO 4
CALL RESET
NMEAS=PARM(1)+.5
NSIM=PARM(2)+.5
NOUT=PARM(3)+.5
MSEND=PARM(4)+.5
GO TO 1
CONTINUE
4 IF(MODEL(1) .NE. PRIN) GO TO 5
CALL DMPOUT(NOUT)
GO TO 1
CONTINUE
5 IF(MODEL(1) .EQ. FINI) CALL FINISH(NOUT)
IF(MODEL(1) .NE. NEWL) GO TO 8
CALL START(AMDA, NO, ISEED)
GO TO 1
CONTINUE
DO 7 K=1,300
HOLDX(K)=0.0D+00
HOLDXX(K)=0.0D+00
7 CONTINUE
8 CONTINUE
NEND=NO*.6
NSEED=IABS(ISEED)/2*2+1
DO 6 K=1,NSIM
CALL SML((NSEED, FITTED, ON)
IF(MODEL(1) .EQ. WOOD) CALL WOODS(AMDA, NO, FITTED)
IF(MODEL(1) .EQ. GENE) CALL GENES(AMDA, NO, FITTED)
IF(MODEL(1) .EQ. WEIS) CALL WEISS(AMDA, NO, FITTED)
IF(MODEL(1) .EQ. WOLM) CALL WOLMAN(AMDA, NO, PARM(1), FITTED)
DO 9 LMN=1,NEND
HOLDX((LMN)=FITTED((LMN)+HOLDX(LMN)
HOLDXX(LMN)=HOLDXX((LMN)+FITTED((LMN)*2
CONTINUE
6 IF(ON) WRITE(NOUT,888) MODEL, NSIM
FORMAT(6,10A4) MODEL, NSIM
100 1/ LAMDA ESTIMATES WILL NOT BE PRINTED FOR EACH SIMULATION. }
ON=.FALSE.
CONTINUE
DO 10 K=1, NEND

```

```

FITTED(K)=HOLDXX(K)/NSIM
HOLDXX(K)=HOLDXX(K)-HOLDX(K)*FITTED(K)
IF(NSIM.GT.1)HOLDXX(K)=HOLDXX(K)/(NSIM-1)
IF(HOLDXX(K).LE.0.0D+00)HOLDXX(K)=0.0D+00
HOLDXX(K)=DSQRT(HOLDXX(K))
CONTINUE
X=COMP(AMDA,NO,FITTED)
WRITE(6,102)
FORMAT(6,102)
102 1,' MEASURES OF PERFORMANCE',/, '0', &-----+
     2,' | SUM OF ERRS | | SUM OF ABS | | SUM OF SQRS |
     3,' | SUM OF RELS | | SUM OF ABS | | SUM OF SQRS |
     4,' MAX ABS ERR | | MAX ABS REL | | MAX ABS REL |
     5,' 8( +-----+), /)
DO 3 L=1,6
M=L
DO 2 K=1,8
MEAS(K,M)=XMEAS(K,L)
2  CONTINUE
FORMAT(6,103)(MEAS(KK,M),KK=1,8),M
103 FORMAT(1X,8(1X,E13.6,1X)),<=EST #,LL)
CONTINUE
CALL PLOTEM(AMDA,NO,FITTED,MODEL)
DO 222 NV=NMEAS,MSEND
CALL STORE(NV,MEAS,MODEL)
CONTINUE
IPNCH=9
FORMAT(1PNCH,777)MODEL,(AMDA(K),FITTED(K),HOLDXX(K),K=1,NEND)
GO TO 1
CONTINUE
FUNCTION COMP(AMDA,NUM,FITTED)
DIMENSION FITTED(NUM,1)
DIMENSION AMDA(NUM,1),ERROR(50,6),REL(50,6)
COMMON/WORK/ERROR,REL
REAL#8 SUM
DO 2 J=1,6
DO 1 I=1,NUM
ERROR(I,J)=AMDA(I,J)-FITTED(I,J)
REL(I,J)=ERROR(I,J)/AMDA(I,J)
CONTINUE
COMPREL(50,6)
RETURN XMEAS(NTYPE,NOEST)
ENTRY XMEAS(NTYPE,NOEST)
GO TO (10,20,30,40,50,60,70,80),NTYPE
CONTINUE
10

```

```

SUM=0.0D+00
DO 11 K=1,NUM
SUM=SUM+ERROR(K,NOEST)
CONTINUE
XMEAS=SUM
RETURN
11
CONTINUE
SUM=0.0D+00
DO 21 K=1,NUM
SUM=SUM+ABS(ERROR(K,NOEST))
CONTINUE
XMEAS=SUM
RETURN
21
CONTINUE
SUM=0.0D+00
DO 31 K=1,NUM
SUM=SUM+ERROR(K,NOEST)**2
CONTINUE
XMEAS=SUM
RETURN
30
CONTINUE
SUM=0.0D+00
DO 41 K=1,NUM
SUM=SUM+REL(K,NOEST)
CONTINUE
XMEAS=SUM
RETURN
40
CONTINUE
SUM=0.0D+00
DO 51 K=1,NUM
SUM=SUM+ABS(REL(K,NOEST))
CONTINUE
XMEAS=SUM
RETURN
50
CONTINUE
SUM=0.0D+00
DO 61 K=1,NUM
SUM=SUM+ABS(REL(K,NOEST))
CONTINUE
XMEAS=SUM
RETURN
60
CONTINUE
SUM=0.0D+00
DO 71 K=1,NUM
SUM=SUM+REL(K,NOEST)**2
CONTINUE
XMEAS=SUM
RETURN
70
CONTINUE
SUM=0.0D+00
DO 71 K=1,NUM
IF(ABS(ERROR(K,NOEST)).GT.SUM)SUM=ABS(ERROR(K,NOEST))
CONTINUE
XMEAS=SUM
RETURN
71

```

```

80      CONTINUE
SUM=0.0D+00
DO 81 K=1,NUM
IF(ABS(.REL(K,NOEST)).GT.SUM) SUM=ABS(.REL(K,NOEST))
CONTINUE
XMEAS=SUM
RETURN
END
SUBROUTINE PLOTEM(ANUM,WORDS)
DIMENSION A(1,1),F(1,1),WORDS(1)
COMMON/WORK/ORDER(600)
DIMENSION TITLE(23)
DIMENSION ARRAY(50,2,7)
DATA TITLE/92HPLOT OF EST NO 1-6(CURVE 1-6) AND TRUE(CURVE 7) FOR
1 DIMENSION INUM(7)
NUM=50
NPLOT=7
DO 1 K=1,7
INUM(K)=NO
CONTINUE
DO 3 I=1,NO
DO 2 K=2,7
KM1=K-1
ARRAY(I,2,KM1)=F(I,KM1)
ARRAY(I,1,KM1)=I
CONTINUE
ARRAY(I,2,7)=A(I,1)
ARRAY(I,1,7)=I
CONTINUE
MINUS=-1
DO 5 L=1,10
LP13=L+13
TITLE(LP13)=WORDS(L)
CONTINUE
CALL PPLOT(NUM,NPLOT,INUM,ARRAY,5,SHLAMDA,10,10HFAILURE NO,
1 MINUS,MINUS,92,TITLE,ORDER)
RETURN
END
SUBROUTINE START(AMDA,NU,IX)
NAMELIST /LAMDA/NUM,NOFMT
NAMELIST /TIMES/NOREP,ISEED
DIMENSION FITTED(1)
COMMON/WORK/FORM(600)
COMMON/EXPTIM/NOR,T(2500)
DIMENSION AMDA(1)
REAL*4 BLANK(3)/0,0,0/
LOGICAL ON

```

```

        FORM(1)=BLANK(1)
        DO 1 K=2,599
        FORM(K)=BLANK(2)
        CONTINUE
1      FORM(600)=BLANK(3)

        NUM=1
        NOFMT=1
        READ(5,LAMDA)
        NTOT=NO FMT*20
        READ(5,100)(FORM(I),I=2,NTOT)
        FORMAT(20A4)
        READ(5,FORM)(AMDA(I),I=1,NUM)
        ISEED=123456789
        NOREP=1
        READ(5,TIMES)
        NCR=NOREP
        NO=NUM
        IX=ISEED
        LDSEED=0
        RETURN
        ENTRY SMLT(INSEED,FITTED,ON)
        IX=NSEED
        IF(IX.EQ.LDSEED) GO TO 25
        LDSEED=IX
        N=NO*NOR
        CALL EXPON(IX ,T,N)
        INDX=0
        DO 2 K=1,NUM
        DO 2 J=1,NOREP
        INDX=INDX+1
        T(INDX)=T(INDX)/AMDA(K)
        CONTINUE
        CALL EST(AMDA,NO NOR)
        IF(ON) WRITE(6,101) NOR
        FORM( /,LAMDA AND LAMDA ESTIMATES ',/,'
        1  EST #1 =>MLE,EST #3 =>MLE(INTEGER DATA),
        2  EST #4 =>RMLE,EST #5 =>RMLE(INTEGER DATA),EST #6 =>ATR
        3  SUBUTE DATA,/,BASED ON,12,REPITITIONS,/)
        CONTINUE
        DO 12 J=1,6
        LL=J*NO
        LL=LL-NO+1
        DO 11 K=L,LL
        FITTED(K)=AMDA(K)
        CONTINUE
        IF(IX.EQ.LDSEED) GO TO 12
        IF( NOT.ON) GO TO 12
        WRITE(6,105) J,(AMDA(K),K=L,LL)
100
101
102
103
104
105
111

```

```

105 FORMAT('OEST #',11,' => ',(T11,8(1X,E13.6,1X)) )
112 CONTINUE
NSEED=IX
RETURN
END
SUBROUTINE EST(AMDA,NUM,T,NOREP)
DIMENSION AMDA(NUM,1) T(NOREP,1)
DOUBLE PRECISION SUM,SUMI
DUMV=ALOG(100.0)
DO 1 K=1,NUM
1   GET TBARS
SUM=0.0D+00
SUMI=0.0D+00
DO 2 J=1,NOREP
SUM=SUM+T(J,K)
T=T(J,K)
2   SUMI=SUMI+T
CONTINUE
GET MAXIMUM LIKELIHOOD ESTIMATORS
V=DUMV
IF(SUM>0.0D+00) V=NOREP/SUM
AMDA(K,2)=V
AMDA(K,4)=V
V=DUMV
IF(SUMI>0.0D+00) V=NOREP/SUMI
VI=1.0D+00+V
VI=ALOG(VI)
AMDA(K,3)=VI
AMDA(K,5)=VI
GET RESTRICTED MLE IF NECESSARY
KM1=K-1
DO 3 I=4,5
IM2=I-2
DO 4 L=1,KM1
LL=K-L
IF(AMDA(LL,I).GE.AMDA(K,I)) GO TO 5
CONTINUE
4   LL=0
CONTINUE
5   LL=LL+1
IF(LL>E-K) GO TO 3
SUM=0.0D+00
DO 6 L=LL,K
TEST=AMDA(L,IM2)
IF(TEST<LT.0.0) TEST=DUMV
SUM=SUM+1.0D+00/TEST
CONTINUE
LEN=K-LL+1

```

```

V=LEN(SUM
1F(1.EQ.5)V=ALOG(1.0+V)
DO 7 L=LL,K
AMDA(L,I)=V
CONTINUE
C ATTRIBUTE DATA
SUMI=0.OD+00
DO 8 J=1,NOREP
IT=T(J,K)
IF(IT.GE.1)SUMI=SUMI+1.0D+00
CONTINUE
IF(SUMI.EQ.0.0D+00)SUMI=01D+00
IF(SUMI.GE.NREP)SUMI=SUMI-1.0D-02
PS=SUMI/NOREP
AMDA(K,6)=-ALOG(PS)
CONTINUE
RETURN
END
SUBROUTINE FINISH(NOUT)
CALL DMPOUT(NOUT)
WRITE(6,100)
STOP
FORMAT(1 PROGRAM TERMINATING DO TO "FINISH" COMMAND OR "/**/")
END
SUBROUTINE WOLMAN(AMDA,NO,B,FITTED)
EXTERNAL RQF
COMMON/EXPTIM/NOR,T(2500)
COMMON/WORK/W(600)
DIMENSION FITTED(NO,1)
REAL*8 R,Q
DIMENSION AMDA(NO,1),X(50),Y(50)
EQUIVALENCE(X(1),R),(X(3),Q)
DUM=ALOG(1.0E+30)
N=NO
B1=B
WRITE(6,201)BI
FORMAT(1/,1,WOLMAN MODEL ESTIMATES,LAMDA(I) = -LN(1-R-Q*(1-
1E13*6,)***(1-1))#/,*EST#2,THROUGH #6 ARE FROM REGRESSION*,/
2,ON LAMDA ESTIMATES #2 THROUGH #6,/*EST#1 IS A MLE BASED ON LAMDA ESTIMATE #3*,//)
3 KSTART=2
DO 1 K=KSTART,6
DO 2 L=1,N
Y(L)=1.0-EXP(-AMDA(L,K))
X(L)=BI**((L-1)
CONTINUE
C
201

```

```

W(600)=5.0
CALL FIT(Y,X,RR,QQ,N)
DO 3 L=1,N
  PONENT=1.0-RR-QQ*X(L)
  IF(PONENT.LE.1.0E-30)GO TO 100
  FITTED(L,K)=-ALOG(PONENT)
  GO TO 3
CONTINUE
3 FITTED(L,K)=DUM
CONTINUE
R=RR
Q=QQ
WRITE(6,202)R,Q,K,(FITTED(L,K),L=1,N)
202 FORMAT(10.13,13.6,13.6)
1 /,EST#1,I1,I=1,(11,8(1X,E13.6,1X)) )
C MAXIMUM LIKELIHOOD ESTIMATOR BASED ON ATTRIBUTE DATA
C=RQFI(B1,N)
CALL ZSYST(RQF,1.0D-04,4,2,X,100,W,IER)
IF(IER.NE.0)GO TO 50
DO 12 L=1,N
  PONENT=1.0-X(1)-X(3)*BI**((L-1)
  IF(PONENT.LT.1.0E-30)GO TO 101
  FITTED(L,I)=-ALOG(PONENT)
  GO TO 12
CONTINUE
12 FITTED(L,I)=DUM
CONTINUE
K=1
WRITE(6,202)R,Q,K,(FITTED(L,K),L=1,N)
GO TO 11
CONTINUE
11 DO 44 L=1,N
  FITTED(L,I)=0.0
  CONTINUE
44 RETURN
END
SUBROUTINE FIT(Y,X,A,B,N)
COMMON/WORK/Y(600)
DIMENSION Y(1),X(1)
DOUBLE PRECISION SUMX,SUMY,SUMXY,SUMXX
DATA IENTER/0/
IENTER=IENTER+1
M=Y(600)+.5
NN=N
SUMX=0.0
SUMY=0.0
PUB02730
PUB02740
PUB02750
PUB02770
PUB02780
PUB02790
PUB02800
PUB02810
PUB02820
PUB02830
PUB02840
PUB02850
PUB02860
PUB02870
PUB02880
PUB02900
PUB02910
PUB02920
PUB02930
PUB02940
PUB02950
PUB02960
PUB02970
PUB02980
PUB03000
PUB03010
PUB03020
PUB03030
PUB03040
PUB03050
PUB03060
PUB03070
PUB03080
PUB03090
PUB03100

```

```

SUMXY=0.0
SUMXX=0.0
FORMAT('1PLOT OF DATA TO BE FITTED',//)
CONTINUE
DO 1 K=1,N
  SUMX=SUMX+X(K)
  SUMY=SUMY+Y(K)
  SUMXY=SUMXY+X(K)*Y(K)
  SUMXX=SUMXX+X(K)*X(K)
CONTINUE
XBAR=SUMX/N
BNUM=SUMXY-XBAR*SUMY
BDEN=SUMXX-XBAR*SUMX
B=BNUM/BDEN
A=SUMY/N-B*XBAR
DO 99 K=1,N
  YY(K)=A+B*X(K)
CONTINUE.EQ.M) IENTER=0
RETURN
END
REAL FUNCTION RQFI*B(N)
REAL*8 SUM,BB,X(2),IT(50)
REAL*8 R,Q
REAL*8 RQF
COMMON/EXPTM/NOR,T(2500)
BB=B
NN=N
INDX=0
DO 1 L=1,NN
  SUM=0.0D+00
  DO 4 K=1,NOR
    INDX=INDX+1
    IT=T(INDX)
    SUM=SUM+IT
    TT(L)=SUM/NOR
    CONTINUE
    ENTRY RQF(X,K)
    SUM=0.0D+00
    R=X(1)
    Q=X(2)
    GO TO (10,20),K
  1 CONTINUE
  DO 2 L=1,NN
    SUM=SUM+IT(L)/(1.0D+00-R-Q*BB***(L-1))
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2      SUM=SUM-1.0D+00/(R+Q*BB**(L-1))
CONTINUE
RQF=SUM
RETURN
DO 3 L=1,N
    SUM=SUM+T(L)/(L-1)*OD+00-R-Q*BB**((L-1))*BB**((L-1)
CONTINUE
RQF=SUM
RETURN
END
SUBROUTINE GE(AMDA,NO,FITTED)
DIMENSION AMDA(NO,1),FITTED(NO,1),X(50),Y(50)
COMMON/WORK/W(600)
COMMON/EXPTIM/NOR,T(2500)
REAL*8 CUMT,SUMREP
WRITE(6,201)
FORMAT(//,*GENERAL ELECTRIC MODEL ESTIMATES, LAMDA(T) = *,
1*C*T*(-ALPHA),/* ALL ESTIMATES ARE FROM REGRESSIONS ON *,
2* CORRESPONDING LAMDA ESTIMATES ,//)
W(600)=6.0
DO 3 NOEST=1,6
CUMT=0.0D+00
INDX=0
DO 1 K=1,NO
    SUMREP=0.0D+00
    DO 2 J=1,NCR
        INDX=INDX+1
        TI=T(INDX)
        TI=TI+1
        GO TO 10,10,20,10,20,30,NOEST
    GO TO 2
CONTINUE
SUMREP=SUMREP+TI
GO TO 2
CONTINUE
SUMREP=SUMREP+IT
GO TO 2
CONTINUE
SUMREP=SUMREP+NOR
J=NOR
CONTINUE
CUMT=CUMT+SUMREP
XK=CUMT
X(K)=-ALOG(XK)
IF(AMDA(K,NOEST).LE.1.0E-60)GO TO 300
Y(K)=ALOG(AMDA(K,NOEST))
CONTINUE

```



```

DELTB=B IFDI(B)/DERIV(B)
B=B-DELTB
IF(ABS(DELTB/B).LT.1.0E-04)GO TO 11
CONTINUE
GO TO 12
CONTINUE
A=HAY(B)
DO 4 K=1,NO
FITTED(K,1)=1.0/A+B/A*X(K)
CONTINUE
4
NOEST=1
WRITE(6,202) A,B,NOEST,(FITTED(L,NOEST),L=1,NO)
RETURN
CONTINUE
BB=B
100=100
CALL ZREALI(DBIFDI,1,D-04,0.D+00,0.D+00,4,1,BB,100,IER)
B=BB
IF(IER.EQ.0)GO TO 11
204
WRITE(6,204)
FORMAT(10ERR IN ESTIMATE #6. SOLUTION TO B FAILED TO CONVERGE.)
DC31 K=1,NO
FITTED(K,NOEST)=0.0
CONTINUE
RETURN
END
FUNCTION BFI(NO)
COMMON/EXPTIM/NDR,T(2500)
REAL*8 SUMI,SUMT,SUMTDI
REAL*8 SUMI
SUMT=0.0D+00
SUMTDI=SUMT
INDEX=0
DO 2 K=1,NO
SUM=0.0D+00
DO 1 J=1,NOR
INDEX=INDEX+1
SUM=SUM+T(INDEX)
CONTINUE
SUM=SUM/NOR
SUMT=SUM+SUM
SUMTDI=SUMTDI+SUM/K
CONTINUE
BFI=SUMT/SUMTDI
RETURN
ENTRY BIFDI(B)
SUM=0.0D+00
SUMI=0.0D+00

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DU 3 K=1,NO
SUM=SUM+1.0/(B+K)
SUMI=SUMI+K/(B+K)
CONTINUE
BIFDI=SUMTDI*SUMTDI*SUMI
RETURN
ENTRY DERIV(B)
SUM=0.0D+00
SUMI=0.0D+00
DC 4 K=1,NO
SUM=SUM+1.0/(B+K)**2
SUMI=SUMI+K/(B+K)**2
CONTINUE
DERIV=SUMTDI*SUMI-SUMT*SUM
RETURN
ENTRY HAY(B)
HAY=(B*SUMTDI+SUMT)/NO
RETURN
END
REAL FUNCTION DBIFDI*8(BB)
REAL*8 BB
B=BB
DBIFDI=BIFDI(B)
RETURN
END
SUBROUTINE WOODS(AMDA,NO,FITTED)
DIMENSION AMDA(NO,1),FITTED(NO,1),X(50),Y(50)
COMMON/WORK/W(600)
COMMON/EXPTIM/NOR,T(2500)
TESTALUG(•99999)
WRITE(6,201)
FORMAT(//,WOODS-CHERNUFF MODEL ESTIMATES, LAMDA(I) = -1,
1LN(1.0-EXP(-ALPHA-BETA**((I-1)))!), ALL ESTIMATES, //)
2 ARE FROM REGRESSIONS ON CORRESPONDING LAMDA ESTIMATES, //)
W(600)=6.0
DO 3 NOEST=1,6
DO 1 K=1,NO
X(K)=K-1
IF(AMDA(K,NOEST)*LE-TEST )GO TO 300
Y(K)=-ALOG(1.0-EXP(-AMDA(K,NOEST)))
CONTINUE
CALL FIT(Y,X,ALPHA,BETA,NO)
DO 2 K=1,NO
ONENT=-ALPHA+BETA*X(K)
IF(ONENT.GE.TEST )ONENT=TEST
FITTED(K,NOEST)=-ALOG(1.0-EXP(ONENT))
CONTINUE
2 WRITE(6,202)ALPHA,BETA,NOEST,(FITTED(L,NOEST),L=1,NO)

```

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